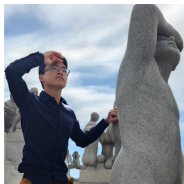


On zeros and algorithms for disordered systems: mean field spin glasses

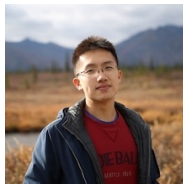
Brice Huang (Stanford → Yale)



Ferenc Bencs
CWI



Daniel Z. Lee
MIT



Kuikui Liu
MIT



Guus Regts
Amsterdam

Ising models

Input: matrix $\mathbf{A} \in \mathbb{R}^{n \times n}$ (**interactions**) and $\beta \geq 0$ (**inv temperature**)

Ising models

Input: matrix $\mathbf{A} \in \mathbb{R}^{n \times n}$ (**interactions**) and $\beta \geq 0$ (**inv temperature**)

Gibbs distribution: prob distribution on $\mathbf{x} \in \{\pm 1\}^n$ with mass

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

for **partition function**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Ising models

Input: matrix $\mathbf{A} \in \mathbb{R}^{n \times n}$ (**interactions**) and $\beta \geq 0$ (**inv temperature**)

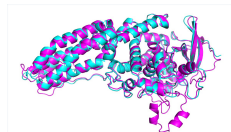
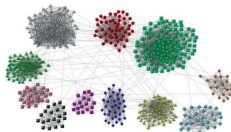
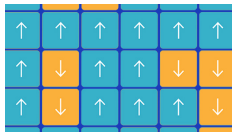
Gibbs distribution: prob distribution on $\mathbf{x} \in \{\pm 1\}^n$ with mass

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

for **partition function**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Applications to physics, statistics, ML, comp bio, ...



Counting and sampling

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A}, \beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Counting and sampling

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A}, \beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Jerrum Valiant Vazirani 86: these are equivalent in many settings

Counting and sampling

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A}, \beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Jerrum Valiant Vazirani 86: these are equivalent in many settings

Main question

When are these problems solvable in poly time, and when are they hard?

Counting and sampling

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A}, \beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Jerrum Valiant Vazirani 86: these are equivalent in many settings

Main question

When are these problems solvable in poly time, and when are they hard?

Key insight: computational complexity related to **phase transitions**

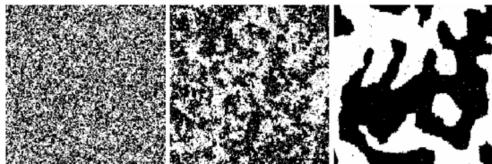


Image from Mora Urbina Cortez Rica 15

Counting and sampling

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

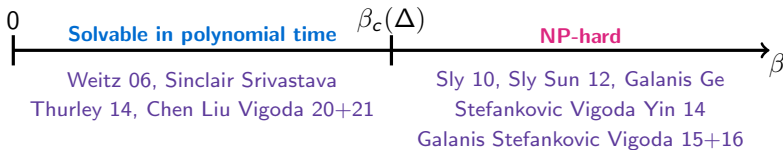
$$\mu_{\mathbf{A}, \beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Jerrum Valiant Vazirani 86: these are equivalent in many settings

Main question

When are these problems solvable in poly time, and when are they hard?

For $\mathbf{A}$ adjacency matrix of graph with max degree Δ :



Beyond worst-case

Main question of this work

What if **A** is random?

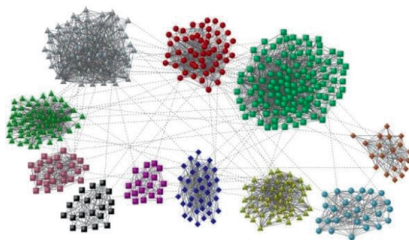
Beyond worst-case

Main question of this work

What if $\mathbf{A}$ is random?

TCS motivation: average case complexity, beyond worst-case analysis

Statistics motivation: Bayesian posterior distributions are random



Physics motivation: models of disordered systems (“spin glasses”)

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

Well-studied in TCS, probability, physics [SK 75](#), [Parisi 79+83](#), [Talagrand 06](#) ...

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

Well-studied in TCS, probability, physics [SK 75](#), [Parisi 79+83](#), [Talagrand 06](#) ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

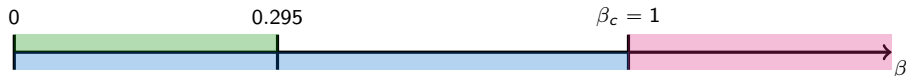
Well-studied in TCS, probability, physics [SK 75](#), [Parisi 79+83](#), [Talagrand 06](#) ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$



Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

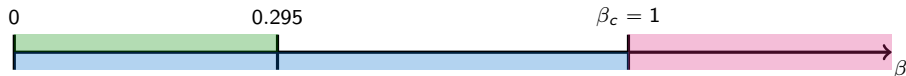
Well-studied in TCS, probability, physics [SK 75](#), [Parisi 79+83](#), [Talagrand 06](#) ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$



MCMC mixes rapidly, samples
to arbitrary accuracy

(Eldan Koehler Zeitouni 19,
Bauerschmidt Bodineau 19,
Anari Jain Koehler Pham
Vuong 22, AKV 24)

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

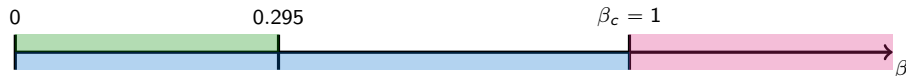
Well-studied in TCS, probability, physics SK 75, Parisi 79+83, Talagrand 06 ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$



MCMC mixes rapidly, samples
to arbitrary accuracy

(Eldan Koehler Zeitouni 19,
Bauerschmidt Bodineau 19,
Anari Jain Koehler Pham
Vuong 22, AKV 24)

(arbitrary accuracy: alg gets desired accuracy ε as input; ε may depend on n)

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

Well-studied in TCS, probability, physics SK 75, Parisi 79+83, Talagrand 06 ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$



MCMC mixes rapidly, samples
to arbitrary accuracy

(Eldan Koehler Zeitouni 19,
Bauerschmidt Bodineau 19,
Anari Jain Koehler Pham
Vuong 22, AKV 24)

Conjectured hard,
“stable” algorithms fail
(El Alaoui Montanari Sellke 22)

(arbitrary accuracy: alg gets desired accuracy ε as input; ε may depend on n)

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

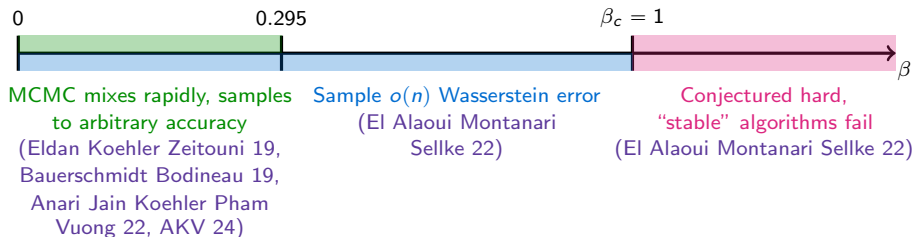
Well-studied in TCS, probability, physics SK 75, Parisi 79+83, Talagrand 06 ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$



(arbitrary accuracy: alg gets desired accuracy ε as input; ε may depend on n)

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

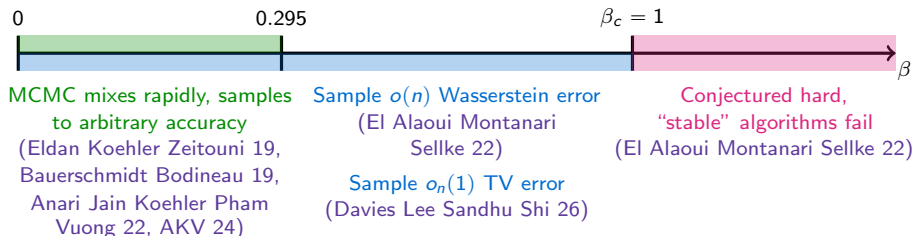
Well-studied in TCS, probability, physics SK 75, Parisi 79+83, Talagrand 06 ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$



(arbitrary accuracy: alg gets desired accuracy ε as input; ε may depend on n)

Sherrington–Kirkpatrick model

SK model: $\mathbf{A} \sim \text{GOE}(n)$ — that is, $A_{i,i} \sim \mathcal{N}(0, \frac{2}{n})$, $A_{i,j} = A_{j,i} \sim \mathcal{N}(0, \frac{1}{n})$

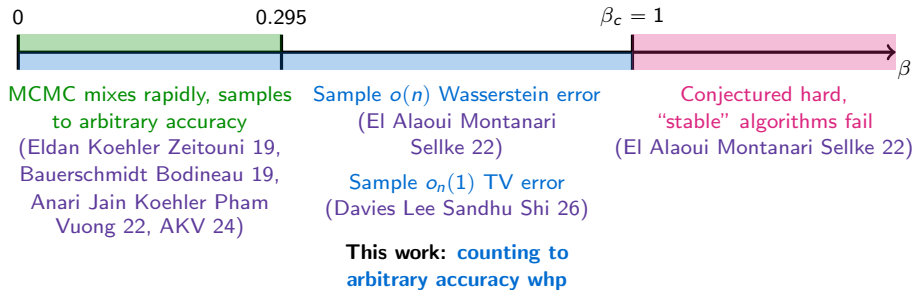
Well-studied in TCS, probability, physics SK 75, Parisi 79+83, Talagrand 06 ...

Counting: estimate **partition fn**

$$Z_{\mathbf{A}}(\beta) = \sum_{\mathbf{x} \in \{\pm 1\}^n} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$

Sampling: sample from **Gibbs dist**

$$\mu_{\mathbf{A},\beta}(\mathbf{x}) = \frac{1}{Z_{\mathbf{A}}(\beta)} \exp\left(\frac{\beta}{2}(\mathbf{x}, \mathbf{A}\mathbf{x})\right)$$



(arbitrary accuracy: alg gets desired accuracy ε as input; ε may depend on n)

Main result

Main theorem for SK model (Bencs H Lee Liu Regts 25)

For any $\varepsilon > 0$, there is an $n^{O(\log(n/\varepsilon))}$ time algorithm which, given $\mathbf{A}$ and $\beta \in [0, 1)$, outputs $\hat{Z}$ such that

$$\mathbb{P}_{\mathbf{A} \sim \text{GOE}(n)} \left[\frac{\hat{Z}}{Z_{\mathbf{A}}(\beta)} \in [1 - \varepsilon, 1 + \varepsilon] \right] = 1 - o_n(1).$$

Main result

Main theorem for SK model (Bencs H Lee Liu Regts 25)

For any $\varepsilon > 0$, there is an $n^{O(\log(n/\varepsilon))}$ time algorithm which, given $\mathbf{A}$ and $\beta \in [0, 1)$, outputs $\hat{Z}$ such that

$$\mathbb{P}_{\mathbf{A} \sim \text{GOE}(n)} \left[\frac{\hat{Z}}{Z_{\mathbf{A}}(\beta)} \in [1 - \varepsilon, 1 + \varepsilon] \right] = 1 - o_n(1).$$

Aizenman Lebowitz Ruelle 87 implies case where ε **independent** of n .
Take $\hat{Z}$ low-degree cluster expansion of $\mathbf{A}$.

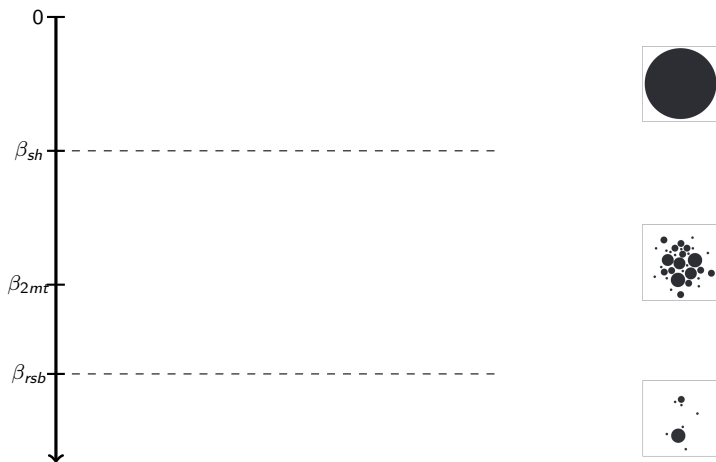
Our result: arbitrary accuracy ε , can depend on n

More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$

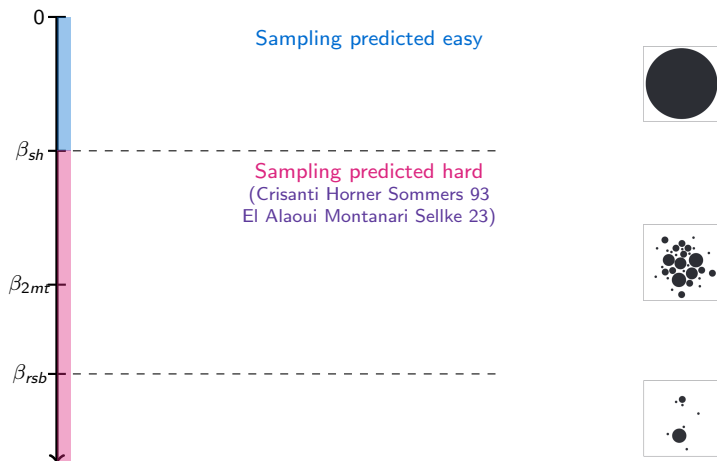
More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$



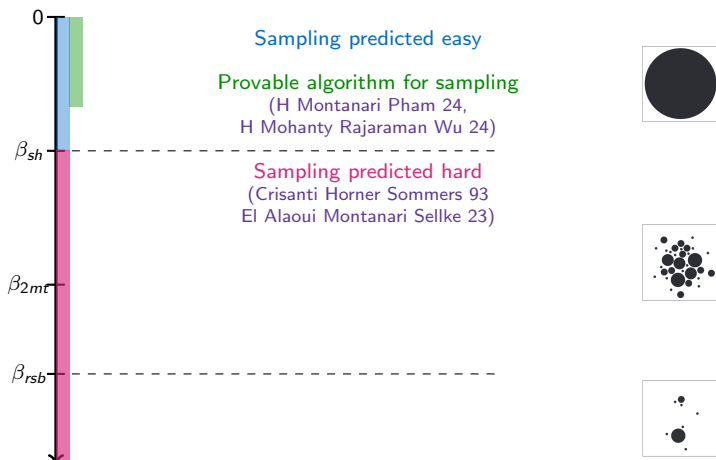
More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$



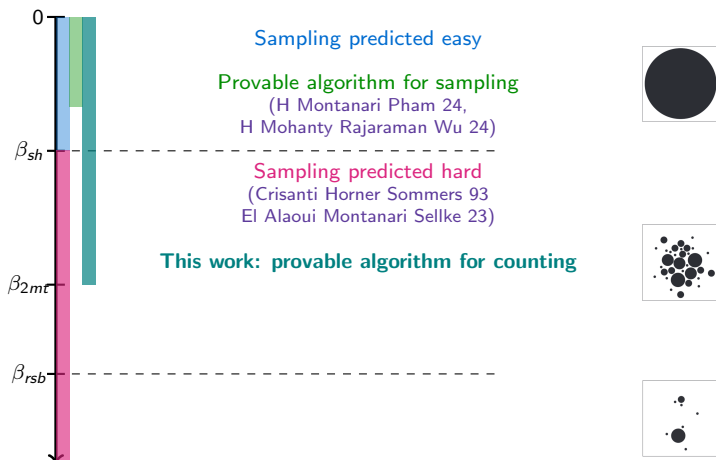
More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$



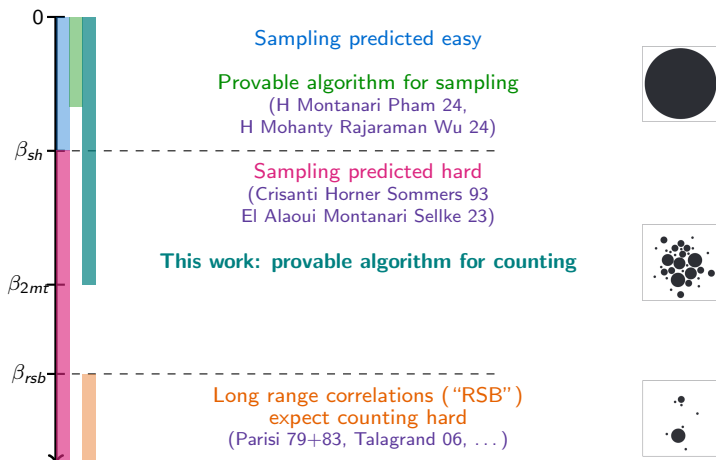
More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$



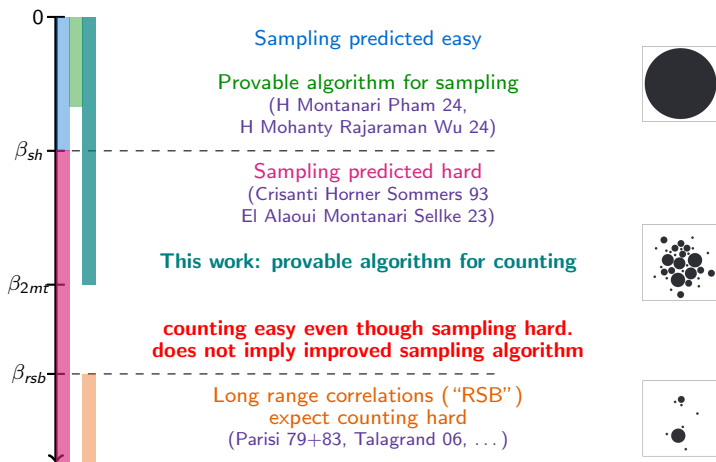
More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$



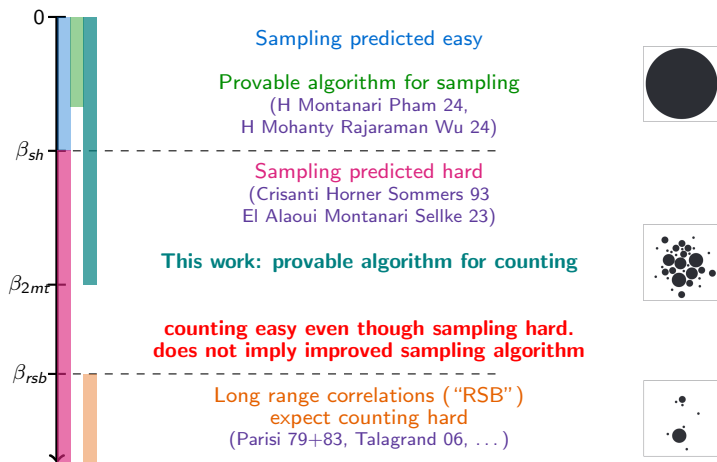
More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$



More generally: (mixed) p -spin models

e.g. 3-spin model: $H(\mathbf{x}) = \sum_{i,j,k=1}^n g_{i,j,k} x_i x_j x_k$, Gibbs distr $\mu(\mathbf{x}) = \frac{1}{Z} e^{\beta H(\mathbf{x})}$



Second moment threshold β_{2mt} : $\max \beta$ where $\mathbb{E}[Z_{\mathbf{A}}(\beta)^2] = O(1) \cdot \mathbb{E}[Z_{\mathbf{A}}(\beta)]^2$

Proof method

Barvinok's interpolation method: if $Z_G(\beta)$ has no zeros in **complex** disk $\mathbb{D}(0, r)$, can estimate $\log Z_G(\beta)$ for $|\beta| < r$ to arbitrary accuracy by truncating its Taylor series

Proof method

Barvinok's interpolation method: if $Z_G(\beta)$ has no zeros in **complex** disk $\mathbb{D}(0, r)$, can estimate $\log Z_G(\beta)$ for $|\beta| < r$ to arbitrary accuracy by truncating its Taylor series

Theorem: zero-freeness of SK model (Bencs H Lee Liu Regts 25)

For any (constant) $r < 1$,

$$\mathbb{P}_{\mathbf{A} \sim \text{GOE}(n)} [Z_{\mathbf{A}}(\beta) \text{ has no zeros in } \mathbb{D}(0, r)] = 1 - o_n(1).$$

Jensen's formula for counting zeros

Jensen's formula

For $R > 0$ and analytic $f : \mathbb{C} \rightarrow \mathbb{C}$ with $f(0) \neq 0$,

$$\sum_{\substack{\omega \in \mathbb{D}(0,R) \\ f(\omega)=0}} \log \left(\frac{R}{|\omega|} \right) = \mathbb{E}_{\theta \sim \text{unif}([0,2\pi])} \left[\log \frac{|f(Re^{i\theta})|}{|f(0)|} \right]$$

Jensen's formula for counting zeros

Jensen's formula

For $R > 0$ and analytic $f : \mathbb{C} \rightarrow \mathbb{C}$ with $f(0) \neq 0$,

$$\sum_{\substack{\omega \in \mathbb{D}(0, R) \\ f(\omega) = 0}} \log \left(\frac{R}{|\omega|} \right) = \mathbb{E}_{\theta \sim \text{unif}([0, 2\pi])} \left[\log \frac{|f(Re^{i\theta})|}{|f(0)|} \right]$$

Set $R = 1 - \frac{\eta}{2}$, $r = 1 - \eta$; $M = \#\{\text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, r)\}$

Jensen's formula for counting zeros

Jensen's formula

For $R > 0$ and analytic $f : \mathbb{C} \rightarrow \mathbb{C}$ with $f(0) \neq 0$,

$$\sum_{\substack{\omega \in \mathbb{D}(0, R) \\ f(\omega) = 0}} \log \left(\frac{R}{|\omega|} \right) = \mathbb{E}_{\theta \sim \text{unif}([0, 2\pi])} \left[\log \frac{|f(Re^{i\theta})|}{|f(0)|} \right]$$

Set $R = 1 - \frac{\eta}{2}$, $r = 1 - \eta$; $M = \#\{\text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, r)\}$

$$\left(\log \frac{R}{r} \right) M \leq \mathbb{E}_{\theta} \left[\log \frac{|Z_{\mathbf{A}}(Re^{i\theta})|}{|Z_{\mathbf{A}}(0)|} \right]$$

Jensen's formula for counting zeros

Jensen's formula

For $R > 0$ and analytic $f : \mathbb{C} \rightarrow \mathbb{C}$ with $f(0) \neq 0$,

$$\sum_{\substack{\omega \in \mathbb{D}(0, R) \\ f(\omega) = 0}} \log \left(\frac{R}{|\omega|} \right) = \mathbb{E}_{\theta \sim \text{unif}([0, 2\pi])} \left[\log \frac{|f(Re^{i\theta})|}{|f(0)|} \right]$$

Set $R = 1 - \frac{\eta}{2}$, $r = 1 - \eta$; $M = \#\{\text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, r)\}$

$$\mathbb{E}_{\mathbf{A}} \left(\log \frac{R}{r} \right) M \leq \mathbb{E}_{\mathbf{A}} \mathbb{E}_{\theta} \left[\log \frac{|Z_{\mathbf{A}}(Re^{i\theta})|}{|Z_{\mathbf{A}}(0)|} \right]$$

Jensen's formula for counting zeros

Jensen's formula

For $R > 0$ and analytic $f : \mathbb{C} \rightarrow \mathbb{C}$ with $f(0) \neq 0$,

$$\sum_{\substack{\omega \in \mathbb{D}(0, R) \\ f(\omega) = 0}} \log \left(\frac{R}{|\omega|} \right) = \mathbb{E}_{\theta \sim \text{unif}([0, 2\pi])} \left[\log \frac{|f(Re^{i\theta})|}{|f(0)|} \right]$$

Set $R = 1 - \frac{\eta}{2}$, $r = 1 - \eta$; $M = \#\{\text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, r)\}$

$$\mathbb{E}_{\mathbf{A}} \left(\log \frac{R}{r} \right) M \leq \frac{1}{2} \mathbb{E}_{\mathbf{A}} \mathbb{E}_{\theta} \left[\log \frac{|Z_{\mathbf{A}}(Re^{i\theta})|^2}{|Z_{\mathbf{A}}(0)|^2} \right]$$

Jensen's formula for counting zeros

Jensen's formula

For $R > 0$ and analytic $f : \mathbb{C} \rightarrow \mathbb{C}$ with $f(0) \neq 0$,

$$\sum_{\substack{\omega \in \mathbb{D}(0,R) \\ f(\omega)=0}} \log \left(\frac{R}{|\omega|} \right) = \mathbb{E}_{\theta \sim \text{unif}([0,2\pi])} \left[\log \frac{|f(Re^{i\theta})|}{|f(0)|} \right]$$

Set $R = 1 - \frac{\eta}{2}$, $r = 1 - \eta$; $M = \#\{\text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, r)\}$

$$\begin{aligned} \mathbb{E}_{\mathbf{A}} \left(\log \frac{R}{r} \right) M &\leq \frac{1}{2} \mathbb{E}_{\mathbf{A}} \mathbb{E}_{\theta} \left[\log \frac{|Z_{\mathbf{A}}(Re^{i\theta})|^2}{|Z_{\mathbf{A}}(0)|^2} \right] \\ &\leq \frac{1}{2} \mathbb{E}_{\theta} \log \left[\mathbb{E}_{\mathbf{A}} \frac{|Z_{\mathbf{A}}(Re^{i\theta})|^2}{|Z_{\mathbf{A}}(0)|^2} \right] \end{aligned}$$

Jensen's formula for counting zeros

Jensen's formula

For $R > 0$ and analytic $f : \mathbb{C} \rightarrow \mathbb{C}$ with $f(0) \neq 0$,

$$\sum_{\substack{\omega \in \mathbb{D}(0,R) \\ f(\omega)=0}} \log \left(\frac{R}{|\omega|} \right) = \mathbb{E}_{\theta \sim \text{unif}([0,2\pi])} \left[\log \frac{|f(Re^{i\theta})|}{|f(0)|} \right]$$

Set $R = 1 - \frac{\eta}{2}$, $r = 1 - \eta$; $M = \#\{\text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, r)\}$

$$\begin{aligned} \mathbb{E}_{\mathbf{A}} \left(\log \frac{R}{r} \right) M &\leq \frac{1}{2} \mathbb{E}_{\mathbf{A}} \mathbb{E}_{\theta} \left[\log \frac{|Z_{\mathbf{A}}(Re^{i\theta})|^2}{|Z_{\mathbf{A}}(0)|^2} \right] \\ &\leq \frac{1}{2} \mathbb{E}_{\theta} \log \left[\mathbb{E}_{\mathbf{A}} \frac{|Z_{\mathbf{A}}(Re^{i\theta})|^2}{|Z_{\mathbf{A}}(0)|^2} \right] \leq O_{\eta}(1) \end{aligned}$$

From $O(1)$ to $o(1)$ zeros

Consequence of Jensen's formula: for all constant $\eta > 0$,

$$\mathbb{E}_{\mathbf{A}} \# \left\{ \text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, 1 - \eta) \right\} = O_{\eta}(1)$$

Not quite enough ...

From $O(1)$ to $o(1)$ zeros

Consequence of Jensen's formula: for all constant $\eta > 0$,

$$\mathbb{E}_{\mathbf{A}} \# \left\{ \text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, 1 - \eta) \right\} = O_{\eta}(1)$$

Not quite enough ...

💡 Re-run the same Jensen's formula strategy with

$$X_{\mathbf{A}}(\beta) = Z_{\mathbf{A}}(\beta) \cdot Y_{\mathbf{A}}(\beta)$$

in place of $Z_{\mathbf{A}}(\beta)$, for carefully chosen **reweight factor** $Y_{\mathbf{A}}(\beta)$

From $O(1)$ to $o(1)$ zeros

Consequence of Jensen's formula: for all constant $\eta > 0$,

$$\mathbb{E}_{\mathbf{A}} \# \left\{ \text{zeros of } Z_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, 1 - \eta) \right\} = O_{\eta}(1)$$

Not quite enough ...

💡 Re-run the same Jensen's formula strategy with

$$X_{\mathbf{A}}(\beta) = Z_{\mathbf{A}}(\beta) \cdot Y_{\mathbf{A}}(\beta)$$

in place of $Z_{\mathbf{A}}(\beta)$, for carefully chosen **reweight factor** $Y_{\mathbf{A}}(\beta)$

$$\Rightarrow \mathbb{E}_{\mathbf{A}} \# \left\{ \text{zeros of } Z_{\mathbf{A}}(\beta) \cdot Y_{\mathbf{A}}(\beta) \text{ in } \mathbb{D}(0, 1 - \eta) \right\} = o_n(1)$$

Conclusion

We give **quasi-polynomial algorithm** to approximate a mean-field spin glass partition function to **arbitrary accuracy**, whp

Conclusion

We give **quasi-polynomial algorithm** to approximate a mean-field spin glass partition function to **arbitrary accuracy**, whp

SK model: works in entire high-temperature regime $\beta < \beta_c = 1$

Mixed p -spin model: works to 2nd moment threshold β_{2mt} .
Surprisingly, counting still possible when sampling hard.

Conclusion

We give **quasi-polynomial algorithm** to approximate a mean-field spin glass partition function to **arbitrary accuracy**, whp

SK model: works in entire high-temperature regime $\beta < \beta_c = 1$

Mixed p -spin model: works to 2nd moment threshold β_{2mt} .
Surprisingly, counting still possible when sampling hard.

Future directions:

- **polynomial-time** algorithm for counting?
- **sparse models**? e.g. hardcore model on random d -reg graph
- stronger evidence of **hardness** beyond β_c ?

Thank you!